

On the Use of Linear Fixed Effects Regression Models for Causal Inference

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Motivation

- Fixed effects models are a primary workhorse for causal inference
- Used for stratified experimental and observational data
- Also used to adjust for **unobservables** in observational studies:
 - ▶ “Good instruments are hard to find ..., so we’d like to have other tools to deal with unobserved confounders. This chapter considers ... strategies that use data with a time or cohort dimension to control for unobserved but fixed omitted variables” (Angrist & Pischke, *Mostly Harmless Econometrics*)
 - ▶ “fixed effects regression can scarcely be faulted for being the bearer of bad tidings” (Green *et al.*, *Dirty Pool*)
- **Question:** What are the exact causal assumptions underlying fixed effects regression models?

Main Methodological Results

- 1 Standard (one-way and two-way) FE estimators are equivalent to particular **matching estimators**
- 2 Identify the information used implicitly to estimate counterfactual outcomes under FE models
- 3 Identify potential sources of bias and inefficiency in FE estimators
- 4 Propose simple ways to improve FE estimators using **weighted** FE regression
- 5 Within-unit matching, first differencing, propensity score weighting, difference-in-differences are all equivalent to weighted FE model with different regression weights
- 6 Offer a specification test for the standard FE model
- 7 Develop fast computation and open-source software

Matching and Regression in Cross-Section Settings

Units	1	2	3	4	5
Treatment status	T	T	C	C	T
Outcome	Y_1	Y_2	Y_3	Y_4	Y_5

- Estimating the Average Treatment Effect (ATE) via matching:

$$Y_1 - \frac{1}{2}(Y_3 + Y_4)$$

$$Y_2 - \frac{1}{2}(Y_3 + Y_4)$$

$$\frac{1}{3}(Y_1 + Y_2 + Y_5) - Y_3$$

$$\frac{1}{3}(Y_1 + Y_2 + Y_5) - Y_4$$

$$Y_5 - \frac{1}{2}(Y_3 + Y_4)$$

Matching Representation of Simple Regression

- Cross-section simple linear regression model:

$$Y_i = \alpha + \beta X_i + \epsilon_i$$

- Binary treatment: $X_i \in \{0, 1\}$
- Equivalent matching estimator:

$$\hat{\beta} = \frac{1}{N} \sum_{i=1}^N \left(\widehat{Y_i(1)} - \widehat{Y_i(0)} \right)$$

where

$$\widehat{Y_i(1)} = \begin{cases} Y_i & \text{if } X_i = 1 \\ \frac{1}{\sum_{i'=1}^N X_{i'}} \sum_{i'=1}^N X_{i'} Y_{i'} & \text{if } X_i = 0 \end{cases}$$
$$\widehat{Y_i(0)} = \begin{cases} \frac{1}{\sum_{i'=1}^N (1-X_{i'})} \sum_{i'=1}^N (1-X_{i'}) Y_{i'} & \text{if } X_i = 1 \\ Y_i & \text{if } X_i = 0 \end{cases}$$

- Treated units matched with the average of non-treated units

One-Way Fixed Effects Regression

- Simple (one-way) FE model:

$$Y_{it} = \alpha_i + \beta X_{it} + \epsilon_{it}$$

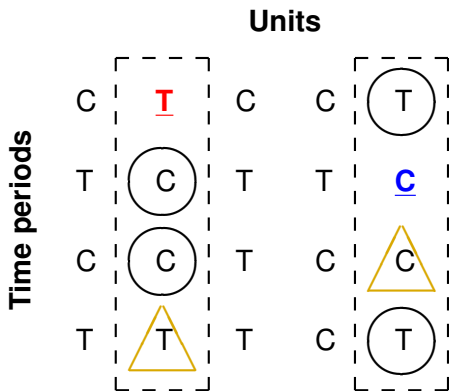
- Commonly used by applied researchers:
 - ▶ **Stratified randomized experiments** (Duflo *et al.* 2007)
 - ▶ **Stratification** and **matching** in observational studies
 - ▶ **Panel data**, both experimental and observational
- $\hat{\beta}_{FE}$ may be biased for the ATE even if X_{it} is exogenous within each unit
- It converges to the weighted average of conditional ATEs:

$$\hat{\beta}_{FE} \xrightarrow{p} \frac{\mathbb{E}\{\text{ATE}_i \sigma_i^2\}}{\mathbb{E}(\sigma_i^2)}$$

where $\sigma_i^2 = \sum_{t=1}^T (X_{it} - \bar{X}_i)^2 / T$

- How are counterfactual outcomes estimated under the FE model?
- Unit fixed effects \implies **within-unit** comparison

Mismatches in One-Way Fixed Effects Model



- T: treated observations
- C: control observations
- **Circles**: Proper matches
- **Triangles**: “Mismatches” \implies attenuation bias

Matching Representation of Fixed Effects Regression

Proposition 1

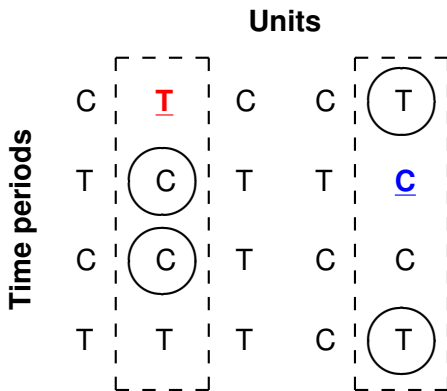
$$\hat{\beta}^{FE} = \frac{1}{K} \left\{ \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \left(\widehat{Y_{it}(1)} - \widehat{Y_{it}(0)} \right) \right\},$$

$$\widehat{Y_{it}(x)} = \begin{cases} Y_{it} & \text{if } X_{it} = x \\ \frac{1}{T-1} \sum_{t' \neq t} Y_{it'} & \text{if } X_{it} = 1 - x \end{cases} \text{ for } x = 0, 1$$

$$K = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \left\{ X_{it} \cdot \frac{1}{T-1} \sum_{t' \neq t} (1 - X_{it'}) + (1 - X_{it}) \cdot \frac{1}{T-1} \sum_{t' \neq t} X_{it'} \right\}.$$

- K : average proportion of proper matches across all observations
- More mismatches \implies larger adjustment
- Adjustment is required except very special cases
- “Fixes” attenuation bias but this adjustment is not sufficient
- Fixed effects estimator is a special case of matching estimators

Unadjusted Matching Estimator



- Consistent if the treatment is exogenous within each unit
- Only equal to fixed effects estimator if heterogeneity in either treatment assignment or treatment effect is non-existent

Unadjusted Matching = **Weighted** FE Estimator

Proposition 2

The unadjusted matching estimator

$$\hat{\beta}^M = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \left(\widehat{Y_{it}(1)} - \widehat{Y_{it}(0)} \right)$$

where

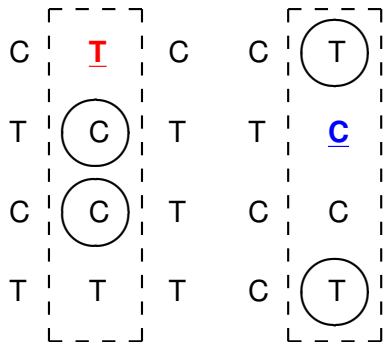
$$\widehat{Y_{it}(1)} = \begin{cases} Y_{it} & \text{if } X_{it} = 1 \\ \frac{\sum_{t'=1}^T X_{it'} Y_{it'}}{\sum_{t'=1}^T X_{it'}} & \text{if } X_{it} = 0 \end{cases} \quad \text{and} \quad \widehat{Y_{it}(0)} = \begin{cases} \frac{\sum_{t'=1}^T (1-X_{it'}) Y_{it'}}{\sum_{t'=1}^T (1-X_{it'})} & \text{if } X_{it} = 1 \\ Y_{it} & \text{if } X_{it} = 0 \end{cases}$$

is equivalent to the weighted fixed effects model

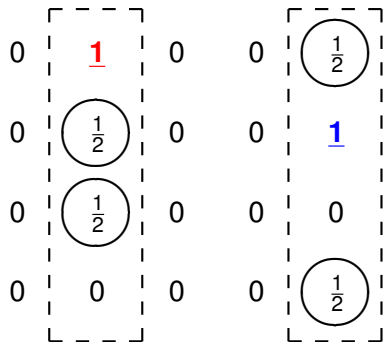
$$\begin{aligned} (\hat{\alpha}^M, \hat{\beta}^M) &= \arg \min_{(\alpha, \beta)} \sum_{i=1}^N \sum_{t=1}^T W_{it} (Y_{it} - \alpha_i - \beta X_{it})^2 \\ W_{it} &\equiv \begin{cases} \frac{T}{\sum_{t'=1}^T X_{it'}} & \text{if } X_{it} = 1, \\ \frac{T}{\sum_{t'=1}^T (1-X_{it'})} & \text{if } X_{it} = 0. \end{cases} \end{aligned}$$

Equal Weights

Treatment



Weights



Different Weights

Treatment				Weights					
C	<u>T</u>	C	C	(T)	0	<u>1</u>	0	0	($\frac{3}{4}$)
T	(C)	T	T	<u>C</u>	0	($\frac{2}{3}$)	0	0	<u>1</u>
C	(C)	T	C	C	0	($\frac{1}{3}$)	0	0	0
T	T	T	C	(T)	0	0	0	0	($\frac{1}{4}$)

- Any within-unit matching estimator leads to weighted fixed effects regression with particular weights
- We derive regression weights given any matching estimator (e.g., first differences) for various quantities (e.g., ATE, ATT)

Fast Computation and Standard Error Calculation

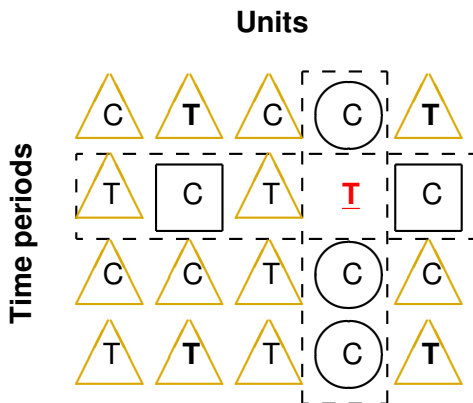
- Standard FE estimator:
 - ▶ “demean” both Y and X
 - ▶ regress demeaned Y on demeaned X
- Weighted FE estimator:
 - ▶ “weighted-demean” both Y and X
 - ▶ regress weighted-demeaned Y on weighted-demeaned X
- Model-based standard error calculation
 - ▶ Various robust sandwich estimators
 - ▶ Easy standard error calculation for matching estimators

Specification Test

- Should we use standard or weighted FE models?
- Standard FE estimator is more efficient if its assumption is correct
- Weighted FE estimator is consistent under the same assumption
- Specification test (White 1980):
 - ▶ Null hypothesis: standard FE model is correct
 - ▶ Does the difference between standard and weighted FE estimators arise by chance?

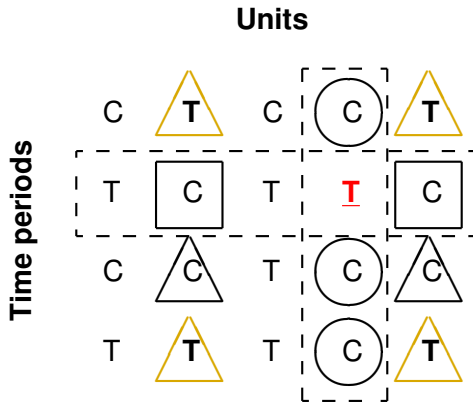
Mismatches in Two-Way FE Model

$$Y_{it} = \alpha_j + \gamma_t + \beta X_{it} + \epsilon_{it}$$



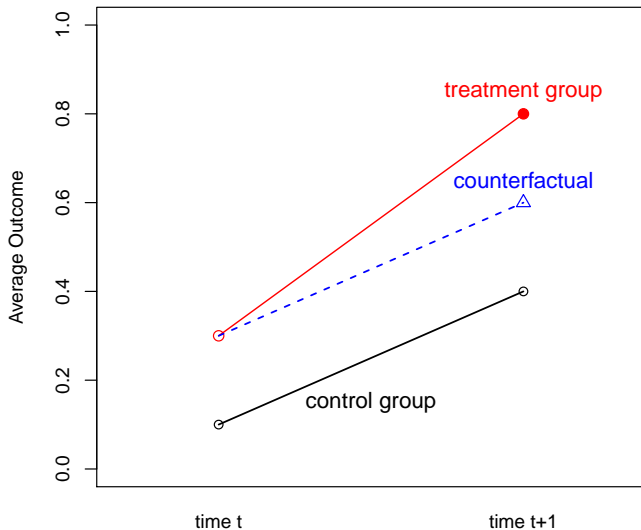
- **Triangles:** Two kinds of mismatches
 - ▶ Same treatment status
 - ▶ Neither same unit nor same time

Mismatches in Weighted Two-Way FE Model



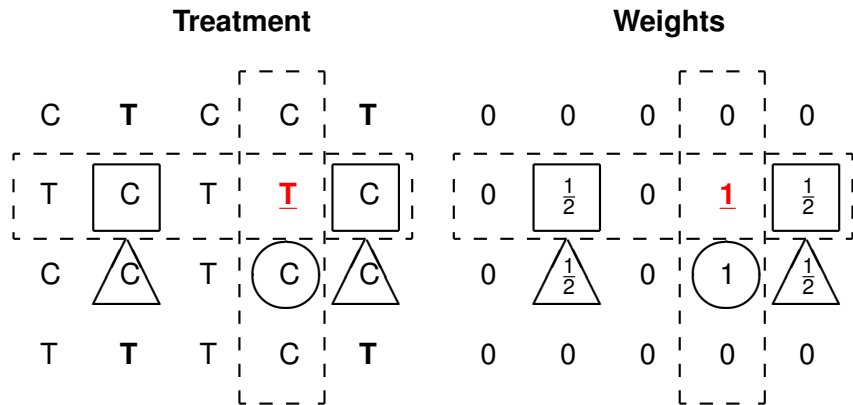
- Some mismatches can be eliminated
- You can NEVER eliminate them all
- Adjustment is always required

Difference-in-Differences (DiD)



General DiD = Weighted Two-Way (Unit and Time) FE

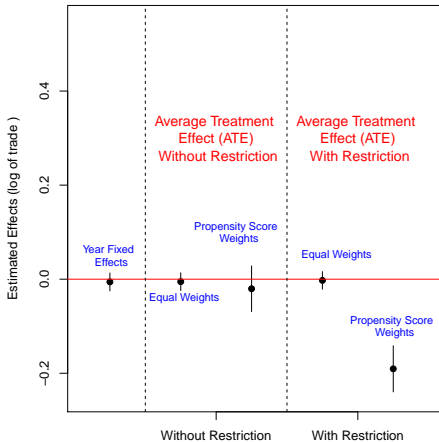
- 2×2 : standard two-way fixed effects
- General setting: Multiple time periods, repeated treatments



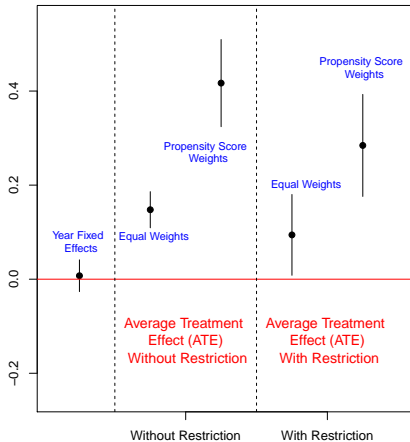
- Weights can be negative \implies the method of moments estimator
- Fast computation via projection on complex plane

Preliminary Empirical Results

(a) Dyad with Both GATT Members
vs.
One GATT Member



(b) Dyad with One GATT Member
vs.
No GATT Member



Concluding Remarks

- Standard one-way and two-way FE estimators are **adjusted** matching estimators
- FE models are not a magic bullet solution to endogeneity
- In many cases, adjustment is not sufficient for removing bias

- Key Question: “Where are the counterfactuals coming from?”
- Different causal assumptions yield different **weighted FE models**
- Weighted FE models encompass a large class of causal assumptions: stratification, first difference, propensity score weighting, difference-in-differences
- Model-based standard error, specification test

- Open source software, R package **wfe**, available